Inferences about Uniqueness in Statistical Learning

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Abstract:
The mind adeptly registers statistical regularities in experience, often incidentally. We used a visual statistical learning paradigm to study incidental learning of predictive relations among animated events. We asked what kinds of statistics participants automatically compute, even when tracking such statistics is task-irrelevant and largely implicit. We find that participants are sensitive to a quantity governing associative learning, ΔP, rather than conditional probabilities or chunk frequencies as previously thought. ΔP specifically reflects the uniqueness, as well as strength, of conditional probabilities. This finding opens the possibility of common, sophisticated inferential mechanisms shared between statistical learning, associative learning, and causal inference scenarios.

Keywords: statistical learning; associative learning

Introduction
A core phenomenon in causal reasoning, contingency learning, and classical conditioning (for a review: Mitchell, De Houwer, & Lovibond, 2009) is that learners do more than register that two stimuli co-occur, but also compute whether they predict each other uniquely and independently, as if attempting to determine a causal model. Suppose two events A and B coincide, such that after most occurrences of A, B occurs. However, B also occurs without A at a very high rate. One would not represent a strong link between A and B in this case. This consideration is captured by a foundational learning formula, ΔP (Allan, 1980; Rescorla & Wagner, 1972; Shanks, 1985):

\[ ΔP = P(\text{A|B}) - P(\text{A|~B}) \]

This equation states that learning is a product of both how often B follows A, as well as how often it appears without it. Surprisingly, it is not known whether this uniqueness principle governs in statistical learning tasks: cases where learning takes place incidentally, below awareness, and in absence of feedback or reward, and in which participants passively observe sequential streams of events (Brady & Oliva, 2008; Kim, Seitz, Feenstra, & Shams, 2009), although it has been reported in a paradigm somewhere in between statistical learning and conditioning (Sobel & Kirkham, 2006).

Experiment
We tested whether learners are sensitive to uniqueness in a visual statistical learning task. Participants saw two distinct event sequences, each composed of a unique set of animated events (Figure 1A), while performing a cover task. Each sequence contained one strongly predictive event pair—a cause and an effect—whose uniqueness we varied. In both sequences, the first term in the ΔP formula (above) was matched: the probability that the effect appeared given the cause appeared on the previous trial was equally high in both. However, in the low ΔP sequence,
we increased the value of the second term, \( \Pr(\text{effect}|\sim\text{cause}) \), by having the effect follow two other events and itself more often than in the high \( \Delta P \) sequence. Thus, the two conditions were matched in terms of the transition probability from cause to effect, as well as in the number of times a cause-effect pair appeared overall (chunk frequency), but differed in terms of how uniquely the cause, rather than other events, predicted the effect. We expected learning to be worse in the low \( \Delta P \) condition.

Following all videos, a surprise forced-choice test probed participants' knowledge of the cause-effect relation in both sequences, separately. The critical questions showed the cause followed by the effect in one video, and the effect followed by the cause in the other; participants had to choose the video that seemed more typical. These questions were matched across conditions in chunk frequency and transition probability. We found that participants were above chance for the high \( \Delta P \) sequence (\( M = 61.83\%, SE = 3.90\% \), \( t(79) = .79, p = .007, \) \( d = 0.31 \)) but below chance for the low \( \Delta P \) sequence (\( M = 41.67\%, SE = 3.89\% \), \( t(79) = -2.16, p = .034, \) \( d = -0.24 \)), which were significantly different from each other (CI [8.43, 29.90], \( t(79) = 3.55, p < .001, d = 0.55 \)), as shown in Figure 2. These differences were due specifically to the difference in \( \Delta P \). Participants had a weaker representation of the cause-effect relationship when uniqueness was low—despite the fact that in both conditions, cause-effect transitions occurred twelve times as often as effect-cause transitions. On the other hand, participants' confidence that they noticed the order among the events was not reliably above 'unsure' for either condition (high \( M = 3.20, SE = 0.13, t(79) = 1.57, p = .121 \); low \( M = 3.06, SE = 0.13, t(79) < 1 \) ), with no difference between them (CI [0.52, 0.91]), Thus, learning was largely implicit, and effects of condition were on the output of this form of learning.

Overall, we conclude that participants' incidental learning is automatically informed by computations of uniqueness, in that neither participants' cover task nor the test questions demanded it or benefited from it. Answers based on chunk frequency or conditional probability were both valid, and computationally simpler, but could not explain the difference in conditions. Thus, participants' incidental learning process can be described as a computation of \( \Delta P \).

Figure 2. Forced-choice test accuracy \( ** = p < .001 \).

**Model**

We developed a computational account of this finding by adapting the Rescorla-Wagner (R-W) learning rule (Rescorla & Wagner, 1972) to the case of sequentially presented stimuli in which the objective is to learn all pairwise strengths among events. To account for the difference in conditions, we required that the weights from all causes to an effect to sum to 1, similar to Bayesian versions of R-W (Kruschke, 2008). This simple normalization step enabled us to capture the difference between conditions, yielding significantly stronger weights for cause-effect than effect-cause links, in both conditions (high \( \Delta P \), cause-effect \( M = 0.70, SE = 0.02 \); effect-cause \( M = 0.11, SE = 0.01, t(79) = 19.22, p < .001 \); low \( \Delta P \), cause-effect \( M = 0.19, SE = 0.01 \), effect-cause \( M = 0.07, SE = 0.01, t(79) = 9.07, p < .001 \), with a significantly larger difference in the high \( \Delta P \) condition (\( t(79) = 17.35, p < .001 \)). Without such normalization, the difference between conditions was not well captured. This is intuitive: the weight from B to X will only be affected by evidence of A to X trials if the weights to X trade off, such that as one gets stronger, the rest weaken (Kruschke, 2008). This can also be seen as representing each link as proportional to the base rate of X, if the base rate of X is, as here, captured in how often it appears following the other events in the state space.

**Conclusion**

Our key finding was that participants in a statistical learning task were sensitive to not only the conditional probability between two events, but also the uniqueness of that relation. This can be seen as the result of normalization: the assumption that predictors of the same effect trade off, and to be considered effective, must raise the probability of the effect above its rate of occurrence otherwise. Overall, these findings bring statistical learning in closer contact with the rich state space.
literature in associative learning and causal reasoning, despite differences in the nature of these learning tasks.

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References


